

Advanced Data Analysis: Principal Component Analysis

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Curse of Dimensionality

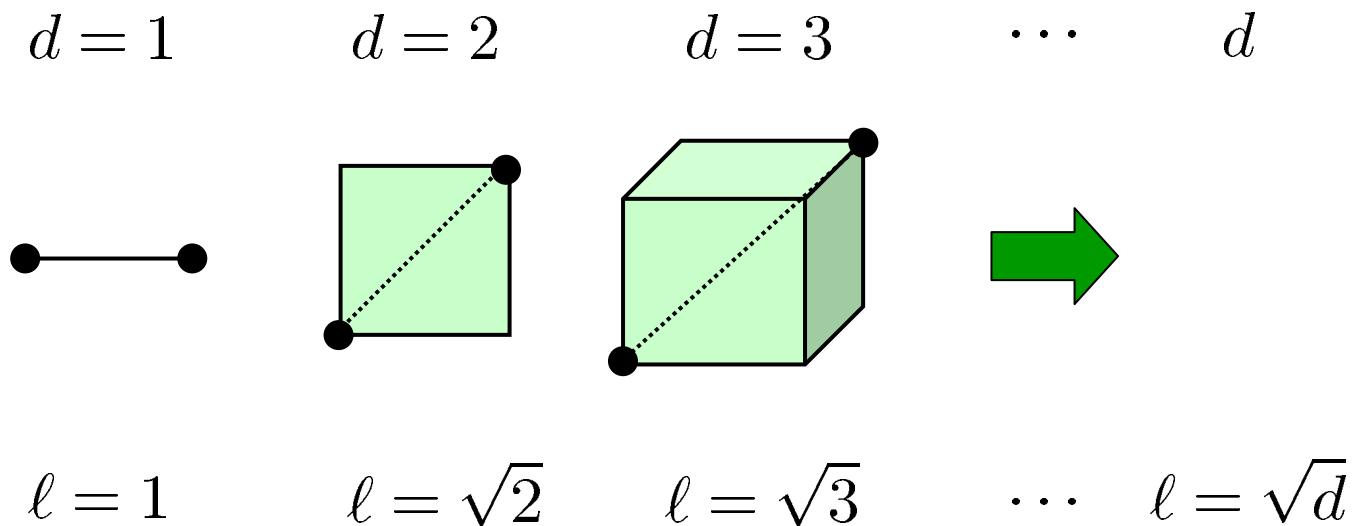
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$$\{\mathbf{x}_i\}_{i=1}^n, \quad \mathbf{x}_i \in \mathbb{R}^d, \quad d \gg 1$$

- If your data samples are high-dimensional, they are often **too complex to directly analyze**.
- Usual geometric intuitions are often only applicable to low-dimensional spaces; such intuitions could be even misleading in high-dimensional spaces.

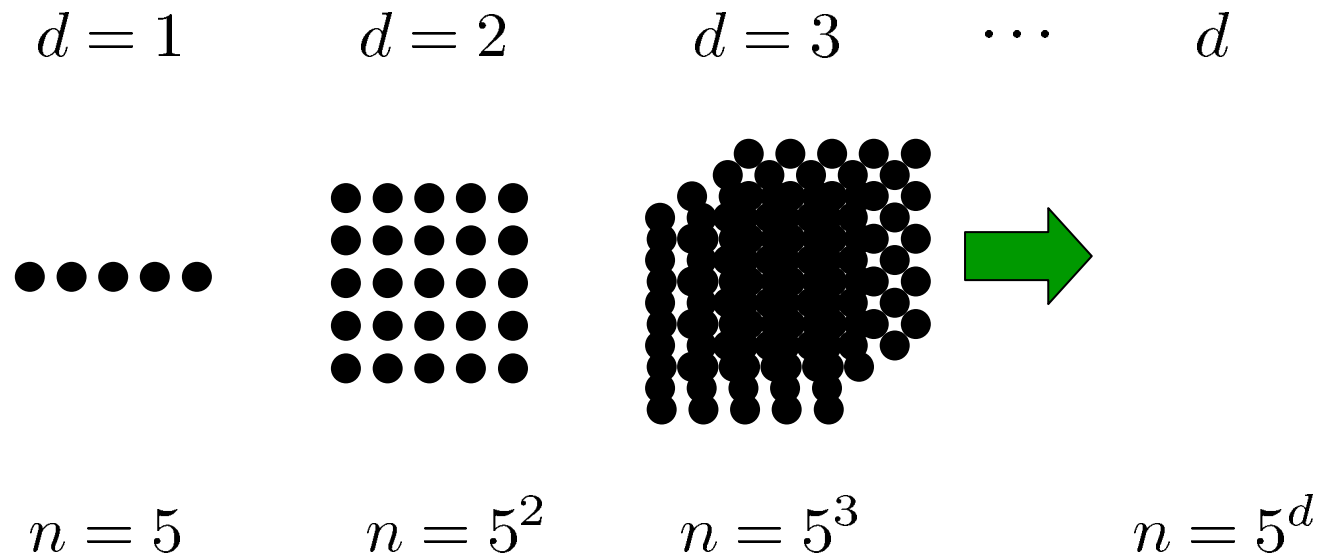
Curse of Dimensionality (cont.) ⁴

- Length of the diagonal line of the unit hypercube gets large as the dimensionality increases.



Curse of Dimensionality (cont.) ⁵

- Grid sampling requires an exponentially large number.



- Unless you have an exponentially large number of samples, your high-dimensional samples are **never dense**.

Dimensionality Reduction

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- We want to reduce the dimensionality of the data while preserving the intrinsic “**information**” in the data.
- Dimensionality reduction is also called **embedding**; if the dimension is reduced up to 3, it is also called **data visualization**.
- **Basic assumption (or belief)** behind dimensionality reduction: your high-dimensional data is redundant in some sense.

Notation: Linear Embedding

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■ Data samples:

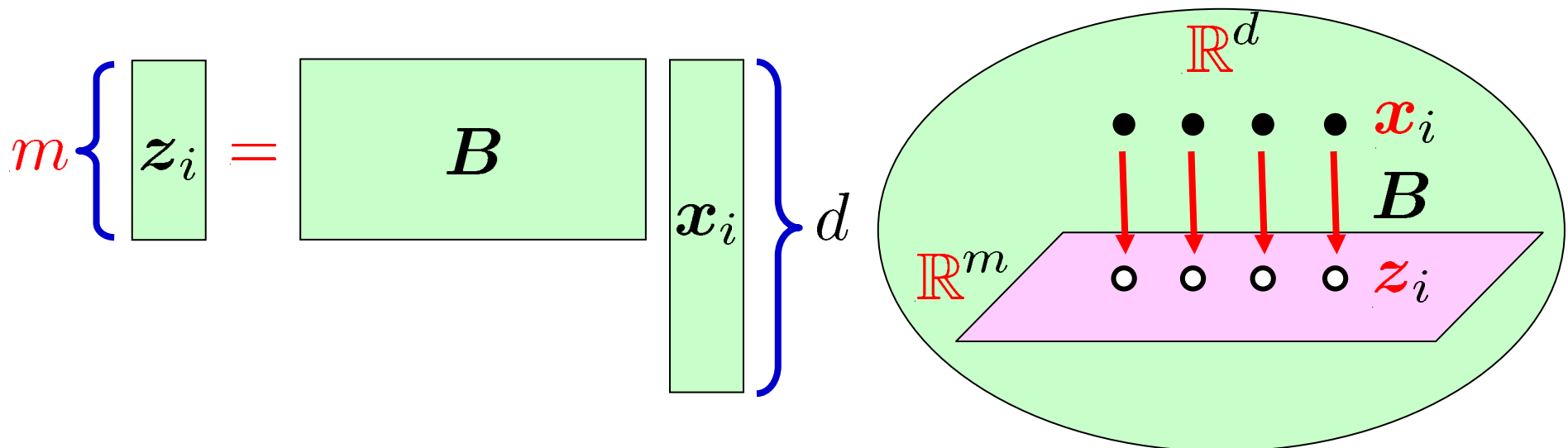
$$\{\mathbf{x}_i\}_{i=1}^n, \quad \mathbf{x}_i \in \mathbb{R}^d, \quad d \gg 1$$

■ Embedding matrix:

$$\mathbf{B} \in \mathbb{R}^{m \times d}, \quad 1 \leq m \ll d$$

■ Embedded data samples:

$$\{\mathbf{z}_i\}_{i=1}^n, \quad \mathbf{z}_i = \mathbf{B}\mathbf{x}_i \in \mathbb{R}^m$$

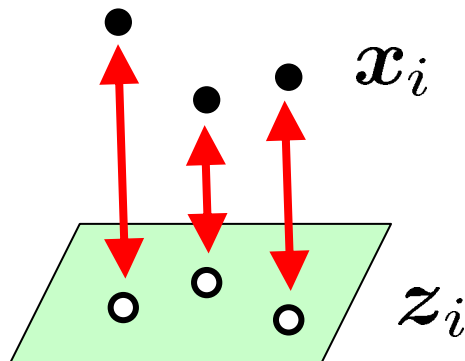


Principal Component Analysis (PCA)⁸

- **Idea:** We want to get rid of a redundant dimension of the data samples

$$\begin{pmatrix} 1 \\ 0 \end{pmatrix}, \begin{pmatrix} 2 \\ 0.1 \end{pmatrix}, \begin{pmatrix} 3 \\ -0.1 \end{pmatrix}$$

- This could be achieved by minimizing the distance between embedded samples and original samples.



Data Centering

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- We **center** the data samples by

$$\bar{x}_i = x_i - \frac{1}{n} \sum_{j=1}^n x_j$$

$$\frac{1}{n} \sum_{i=1}^n \bar{x}_i = 0$$

- In matrix,

$$\bar{X} = XH$$

$$\bar{X} = (\bar{x}_1 | \bar{x}_2 | \cdots | \bar{x}_n)$$

$$X = (x_1 | x_2 | \cdots | x_n)$$

$$H = I_n - \frac{1}{n} \mathbf{1}_{n \times n}$$

I_n : n -dimensional identity matrix

$\mathbf{1}_{n \times n}$: $n \times n$ matrix with all ones

Orthogonal Projection

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- $\{\mathbf{b}_i (\in \mathbb{R}^d)\}_{i=1}^m$: Orthonormal basis in m -dimensional embedding subspace

$$\langle \mathbf{b}_i, \mathbf{b}_j \rangle = \delta_{i,j} = \begin{cases} 1 & (i = j) \\ 0 & (i \neq j) \end{cases}$$

- In matrix, $\mathbf{B}\mathbf{B}^\top = \mathbf{I}_m$

$$\mathbf{B} = (\mathbf{b}_1 | \mathbf{b}_2 | \cdots | \mathbf{b}_m)^\top$$

- Orthogonal projection of $\bar{\mathbf{x}}_i$ is expressed by

$$\sum_{j=1}^m \langle \mathbf{b}_j, \bar{\mathbf{x}}_i \rangle \mathbf{b}_j \quad \left(= \mathbf{B}^\top \mathbf{B} \bar{\mathbf{x}}_i \right)$$

PCA Criterion

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- Minimize the sum of squared distances.

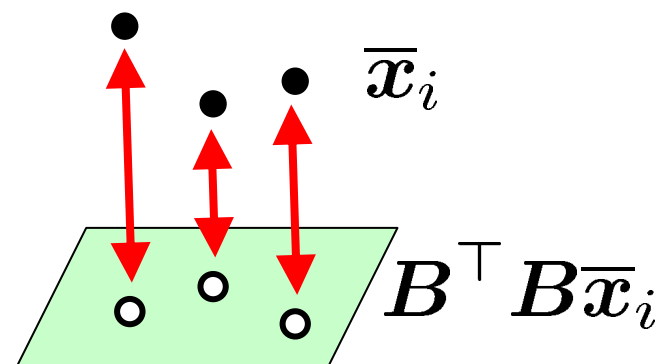
$$\sum_{i=1}^n \|B^\top B \bar{x}_i - \bar{x}_i\|^2 \quad \left(= -\text{tr}(B \bar{C} B^\top) + \text{tr}(\bar{C}) \right)$$

$$\bar{C} = \sum_{i=1}^n \bar{x}_i \bar{x}_i^\top = \bar{X} \bar{X}^\top$$

- PCA criterion:

$$B_{PCA} = \underset{B \in \mathbb{R}^{m \times d}}{\text{argmax}} \text{tr}(B \bar{C} B^\top)$$

$$\text{subject to } B B^\top = I_m$$



PCA: Summary

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■ A PCA solution:

$$\mathbf{B}_{PCA} = (\psi_1 | \psi_2 | \cdots | \psi_m)^\top$$

$\{\lambda_i, \psi_i\}_{i=1}^m$: Sorted eigenvalues and normalized eigenvectors of $\overline{\mathbf{C}}\psi = \lambda\psi$

$$\lambda_1 \geq \lambda_2 \geq \cdots \geq \lambda_d$$

$$\langle \psi_i, \psi_j \rangle = \delta_{i,j}$$

■ PCA embedding of a sample $\mathbf{x}$:

$$\overline{\mathbf{z}} = \mathbf{B}_{PCA} \left(\mathbf{x} - \frac{1}{n} \mathbf{X} \mathbf{1}_n \right)$$

$\mathbf{1}_n$: n -dimensional vector with all ones

■ Lagrangian:

$$L(\mathbf{B}, \mathbf{\Delta}) = \text{tr}(\mathbf{B}\overline{\mathbf{C}}\mathbf{B}^\top) - \text{tr}((\mathbf{B}\mathbf{B}^\top - \mathbf{I}_m)\mathbf{\Delta})$$

$\mathbf{\Delta}$: Lagrange multipliers (symmetric)

■ Stationary point (necessary condition):

$$\bullet \frac{\partial L}{\partial \mathbf{B}} = 2\mathbf{B}\overline{\mathbf{C}} - 2\mathbf{\Delta}\mathbf{B} = 0$$

$$\longrightarrow \overline{\mathbf{C}}\mathbf{B}^\top = \mathbf{B}^\top \mathbf{\Delta} \quad (1)$$

$$\bullet \frac{\partial L}{\partial \mathbf{\Delta}} = \mathbf{B}\mathbf{B}^\top - \mathbf{I}_m = 0$$

$$\longrightarrow \mathbf{B}\mathbf{B}^\top = \mathbf{I}_m \quad (2)$$

■ Eigendecomposition:

$$\mathbf{\Delta} = \mathbf{T}\mathbf{\Gamma}\mathbf{T}^\top \quad (3)$$

$\mathbf{T}$: orthogonal matrix
 $\mathbf{\Gamma}$: diagonal matrix

$$\mathbf{T}^{-1} = \mathbf{T}^\top$$

Proof (cont.)

■ (1) & (3) $\longrightarrow \bar{C}B^\top = B^\top T \Gamma T^\top$ (4)

$\longrightarrow \bar{C}B^\top T = B^\top T \Gamma$

$\longrightarrow \bar{C}F = F\Gamma$ (5) $F = B^\top T$

■ (5) is an eigensystem

$\longrightarrow \mathcal{R}(F) = \text{span}(\{\psi_{k_i}\}_{i=1}^m)$ (6)

$\Gamma = \text{diag}(\lambda_{k_1}, \lambda_{k_2}, \dots, \lambda_{k_m})$ (7)

$k_i \in \{1, 2, \dots, d\}$

■ $\mathcal{R}(F) = \mathcal{R}(B^\top T) = \mathcal{R}(B^\top)$ (8)

■ (6) & (8) $\longrightarrow \mathcal{R}(B^\top) = \text{span}(\{\psi_{k_i}\}_{i=1}^m)$

■ Solution is expressed as

$$B = (\psi_{k_1} | \psi_{k_2} | \dots | \psi_{k_m})^\top$$

Proof (cont.)

- (2) ➡ $\text{rank}(\mathbf{B}) = m$
➡ all $\{k_i\}_{i=1}^m$ are distinct
- We should choose the best $\{k_i\}_{i=1}^m$ that maximizes $\text{tr}(\mathbf{B}\overline{\mathbf{C}}\mathbf{B}^\top)$.
- (4) & (7) ➡

$$\begin{aligned}
 \text{tr}(\mathbf{B}\overline{\mathbf{C}}\mathbf{B}^\top) &= \text{tr}(\mathbf{B}\mathbf{B}^\top \mathbf{T}\mathbf{T}^\top) \\
 &= \text{tr}(\mathbf{T}\mathbf{T}^\top) \\
 &= \text{tr}(\mathbf{T}^\top \mathbf{T}) \\
 &= \sum_{i=1}^m \lambda_{k_i}
 \end{aligned}$$
- $\lambda_1 \geq \lambda_2 \geq \dots \geq \lambda_d$
➡ $k_i = i$ gives a solution.

(Q.E.D.)

Correlation

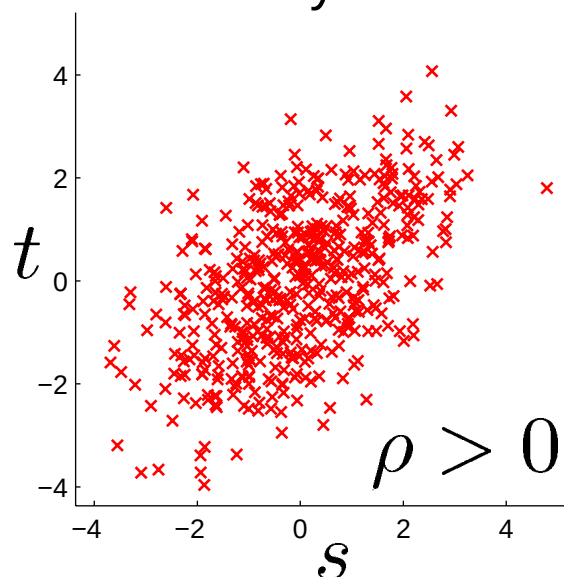
■ Correlation coefficient for $\{s_i, t_i\}_{i=1}^n$:

$$\rho = \frac{\sum_{i=1}^n (s_i - \bar{s})(t_i - \bar{t})}{\sqrt{(\sum_{i=1}^n (s_i - \bar{s})^2) (\sum_{i=1}^n (t_i - \bar{t})^2)}}$$

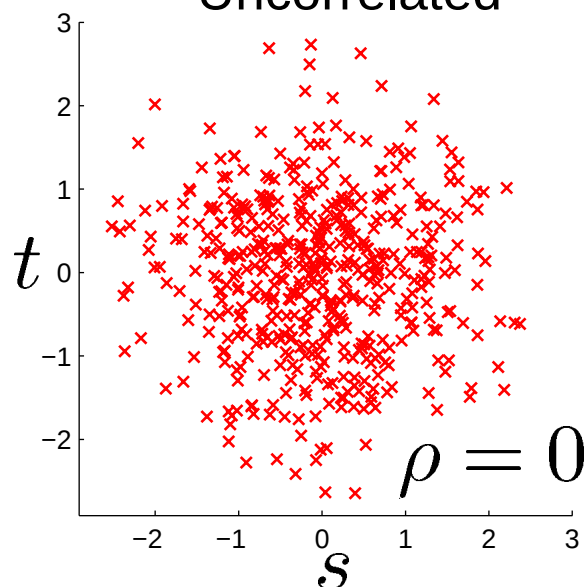
$$\bar{s} = \frac{1}{n} \sum_{i=1}^n s_i$$

$$\bar{t} = \frac{1}{n} \sum_{i=1}^n t_i$$

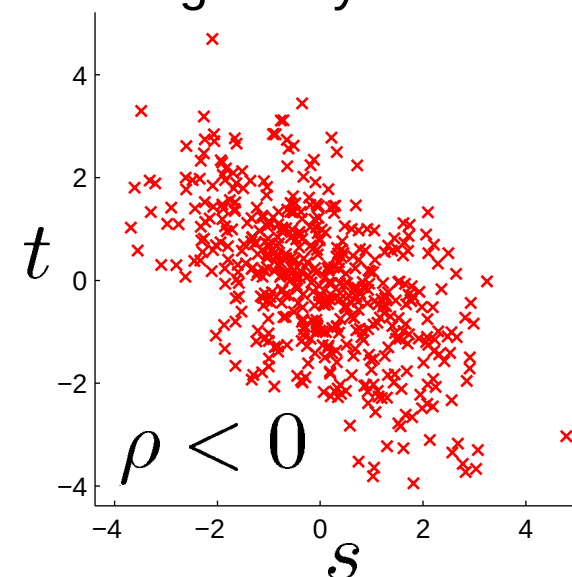
Positively correlated



Uncorrelated



Negatively correlated



PCA Uncorrelates Data

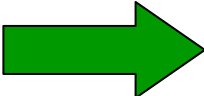
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$$\mathbf{B}_{PCA} = (\boldsymbol{\psi}_1 | \boldsymbol{\psi}_2 | \cdots | \boldsymbol{\psi}_m)^\top$$

- Covariance matrix of the PCA-embedded samples is diagonal.

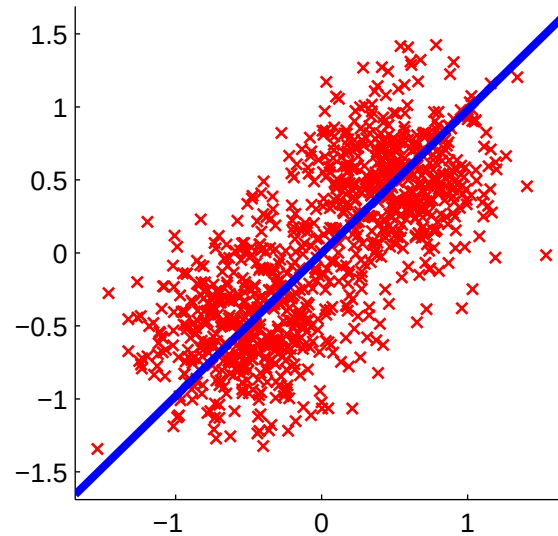
$$\frac{1}{n} \sum_{i=1}^n \bar{\mathbf{z}}_i \bar{\mathbf{z}}_i^\top = \text{diag}(\lambda_1, \lambda_2, \dots, \lambda_m)$$

(Homework)

 Each element in $\bar{\mathbf{z}}$ is uncorrelated!

Examples

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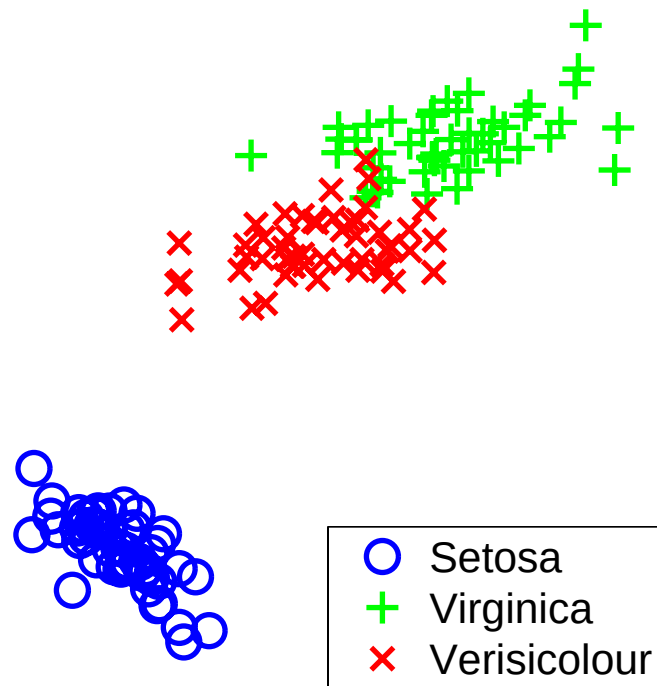


- Data is well described
- PCA is intuitive, easy to implement, analytic solution available, and fast.

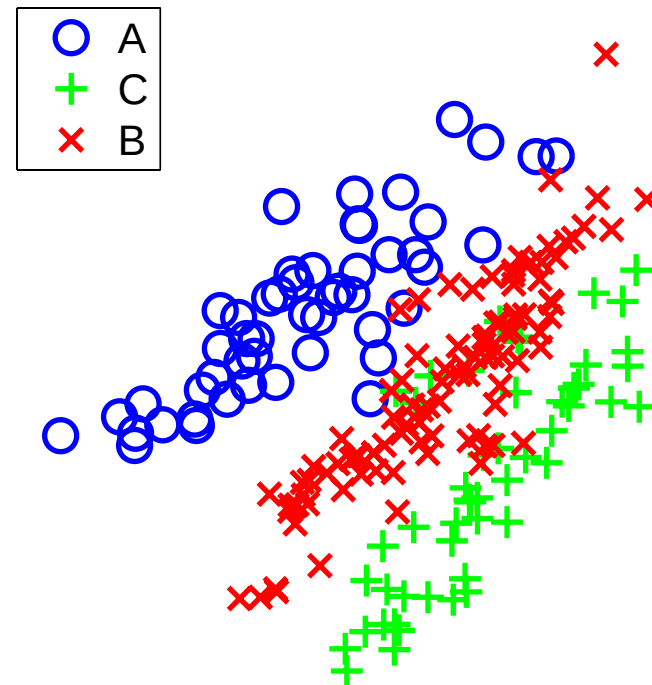
Examples (cont.)

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Iris data (4d->2d)



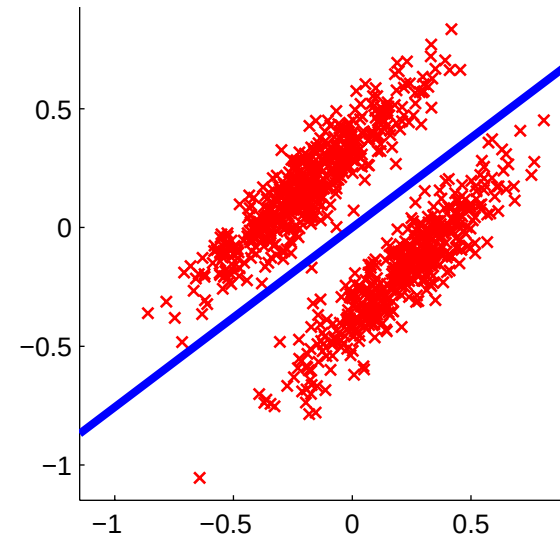
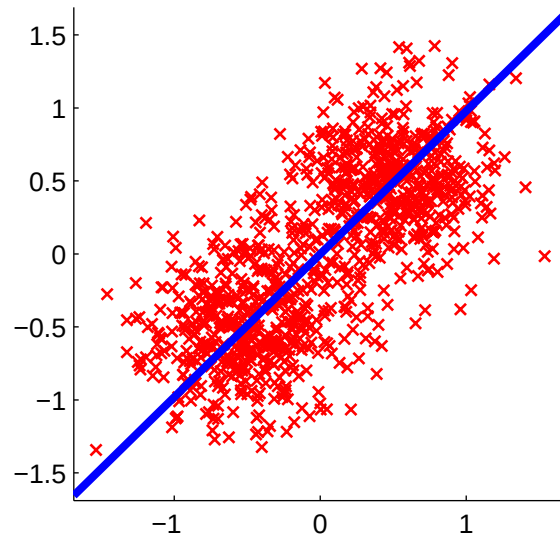
Letter data (16d->2d)



■ Embedded samples seem informative.

Examples (cont.)

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- However, PCA does not necessarily preserve interesting information such as clusters.

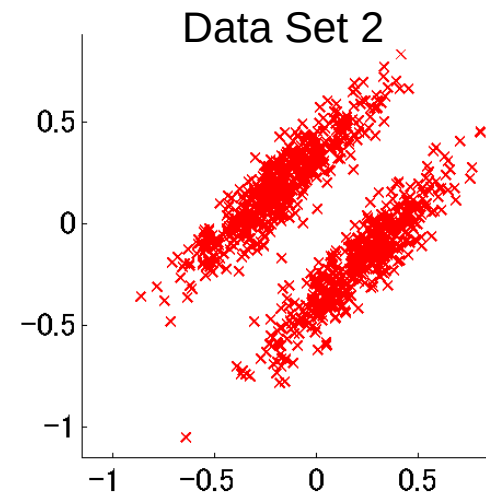
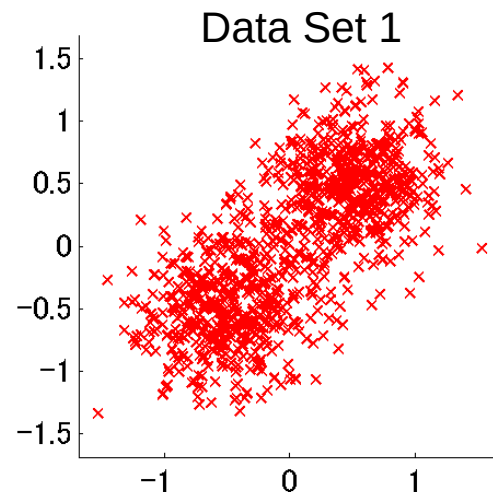
Homework

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1. Implement PCA and reproduce the 2-dimensional examples shown in the class.

- Data sets are available from

<http://sugiyama-www.cs.titech.ac.jp/~sugi/data/DataAnalysis>



- You may create similar (or more interesting) data sets by yourself.

Homework (cont.)

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2. Let

- $B : m \times d, (1 \leq m \leq d)$
- $C, D : d \times d$, positive definite, symmetric
- $\{\lambda_i, \psi_i\}_{i=1}^m$: Sorted **generalized** eigenvalues and normalized eigenvectors of $C\psi = \lambda D\psi$

$$\lambda_1 \geq \lambda_2 \geq \cdots \geq \lambda_d$$

$$\langle D\psi_i, \psi_j \rangle = \delta_{i,j}$$

Prove that a solution of

$$B_{min} = \operatorname{argmin}_{B \in \mathbb{R}^{m \times d}} \left[\operatorname{tr}(BCB^T) \right]$$

$$\text{subject to } BDB^T = I_m$$

is given by

$$B_{min} = (\psi_d | \psi_{d-1} | \cdots | \psi_{d-m+1})^T$$

Homework (cont.)

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3. Prove that PCA uncorrelates the samples; more specifically, prove that the covariance matrix of the PCA-embedded samples is the following diagonal matrix:

$$\frac{1}{n} \sum_{i=1}^n \bar{\mathbf{z}}_i \bar{\mathbf{z}}_i^\top = \text{diag}(\lambda_1, \lambda_2, \dots, \lambda_m)$$

$$\bar{\mathbf{z}}_i = \mathbf{B}_{PCA} \bar{\mathbf{x}}_i$$

$$\mathbf{B}_{PCA} = (\boldsymbol{\psi}_1 | \boldsymbol{\psi}_2 | \dots | \boldsymbol{\psi}_m)^\top$$

Suggestion

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- Read the following article for upcoming classes:
 - X. He & P. Niyogi: Locality preserving projections, In *Advances in Neural Information Processing Systems 16*, MIT Press, Cambridge, MA, 2004.

http://books.nips.cc/papers/files/nips16/NIPS2003_AA20.pdf